

## 4.3 The Normal Distribution

### DEFINITION

A continuous rv  $X$  is said to have a **normal distribution** with parameters  $\mu$  and  $\sigma$  (or  $\mu$  and  $\sigma^2$ ), where  $-\infty < \mu < \infty$  and  $0 < \sigma$ , if the pdf of  $X$  is

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi} \sigma} e^{-(x-\mu)^2/(2\sigma^2)} \quad -\infty < x < \infty \quad (4.3)$$

$e \approx 2.71828$ ,  $\pi \approx 3.14159$ .

**Notation:**  $N(\mu, \sigma^2)$

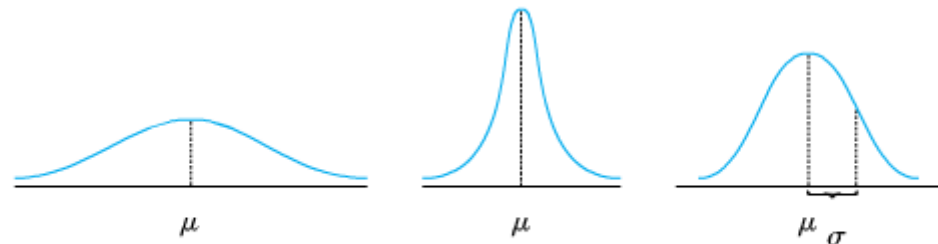


Figure 4.13 Normal density curves

**Theorem:** If  $X$  has normal distribution with parameters  $\mu$  and  $\sigma$  then

$$\begin{aligned} E(X) &= \mu \\ V(X) &= \sigma^2 \end{aligned}$$

### COMPUTATION

**TI-83:** TI-83 has normal probabilities and percentiles programmed in DISTR menu ([2<sup>ND</sup>] [VARS])

$P(a \leq X \leq b) = \mathbf{normalcdf(a, b, \mu, \sigma)}$  ( use  $a = -10^{99}$  if  $a = -\infty$ , and  $b = 10^{99}$  if  $b = \infty$ )

100p%-th percentile =  $\mathbf{invNorm(p, \mu, \sigma)}$

OR Table A3

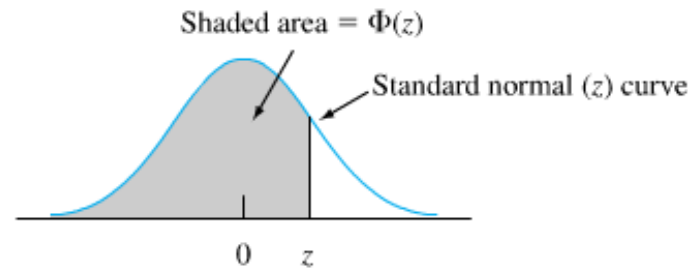
## DEFINITION

The normal distribution with parameter values  $\mu = 0$  and  $\sigma = 1$  is called the **standard normal distribution**. A random variable having a standard normal distribution is called a **standard normal random variable** and will be denoted by  $Z$ . The pdf of  $Z$  is

$$f(z; 0, 1) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad -\infty < z < \infty$$

The graph of  $f(z; 0, 1)$  is called the *standard normal* (or  $z$ ) curve. Its inflection points are at 1 and  $-1$ . The cdf of  $Z$  is  $P(Z \leq z) = \int_{-\infty}^z f(y; 0, 1) dy$ , which we will denote by  $\Phi(z)$ .

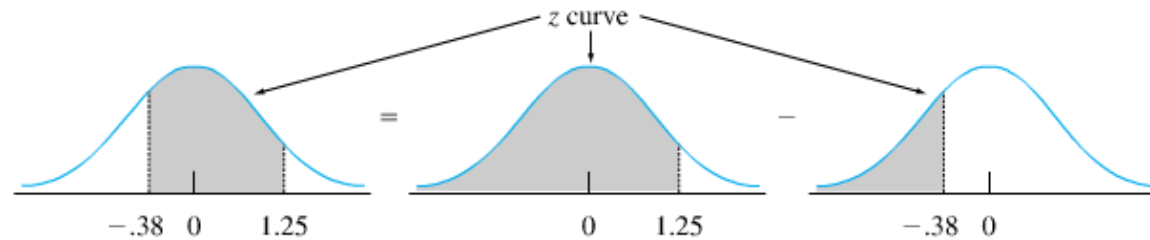
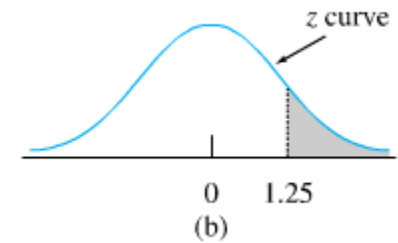
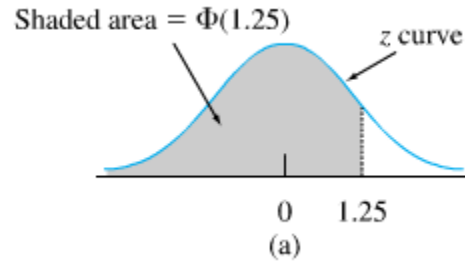
Notation:  $N(0,1)$



**Figure 4.14** Standard normal cumulative areas tabulated in Appendix Table A.3

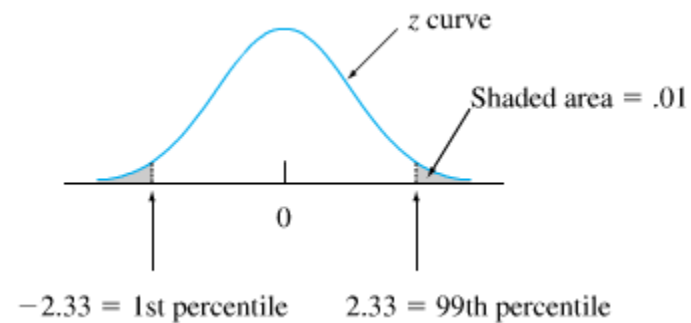
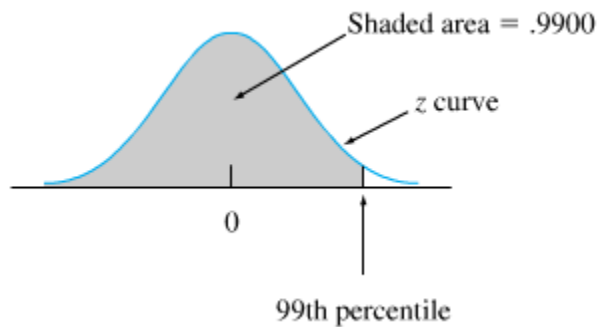
**Example 4.13** Let's determine the following standard normal probabilities: (a)  $P(Z \leq 1.25)$ , (b)  $P(Z > 1.25)$ , (c)  $P(Z \leq -1.25)$ , and (d)  $P(-.38 \leq Z \leq 1.25)$ .

- (a)  $P(Z \leq 1.25) = \Phi(1.25) = 0.8944$
- (b)  $P(Z > 1.25) = 1 - \Phi(1.25) = 0.1056$
- (c)  $P(Z \leq -1.25) = \Phi(-1.25) = 0.1056$
- (d)  $P(-.38 \leq Z \leq 1.25) = \Phi(1.25) - \Phi(-.38)$



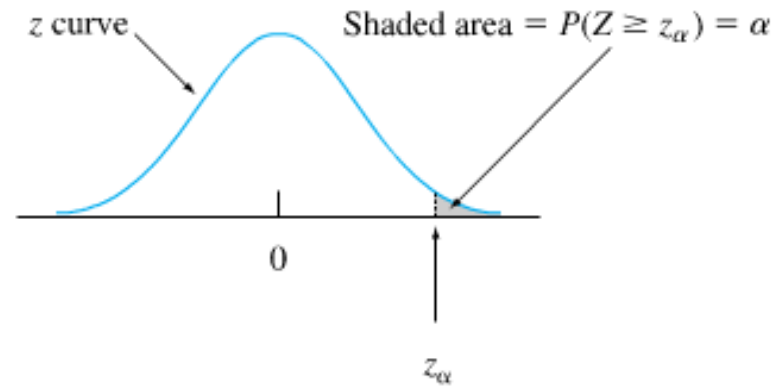
### Finding percentiles of the Standard Normal Distribution:

Example 2. Find the 99<sup>th</sup> percentile and the 1<sup>st</sup> percentile of the standard normal distribution.



Answer: 99<sup>th</sup> percentile  $\approx 2.33$ , 1<sup>st</sup> percentile  $\approx -2.33$

**$z_\alpha$  Value (critical value):**  $z_\alpha$  is the value of Z for which the area under the z-curve and to the right of  $z_\alpha$  is  $\alpha$

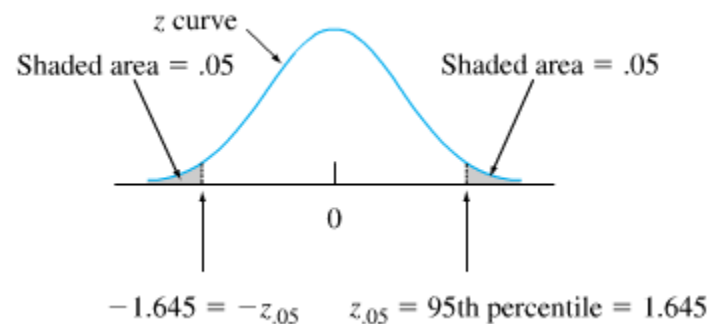


$z_\alpha = (1-\alpha)100\%$  percentile of the standard normal curve

**Table 4.1 Standard Normal Percentiles and Critical Values**

Percentile	90	95	97.5	99	99.5	99.9	99.95
$\alpha$ (tail area)	.1	.05	.025	.01	.005	.001	.0005
$z_\alpha = 100(1 - \alpha)$ th percentile	1.28	1.645	1.96	2.33	2.58	3.08	3.27

Example 3. Find  $z_{0.05}$ . Answer:  $z_{0.05} = 95^{\text{th}}$  percentile of standardnormal distribution = 1.645



## Relation between nonstandard and standard normal distributions:

### PROPOSITION

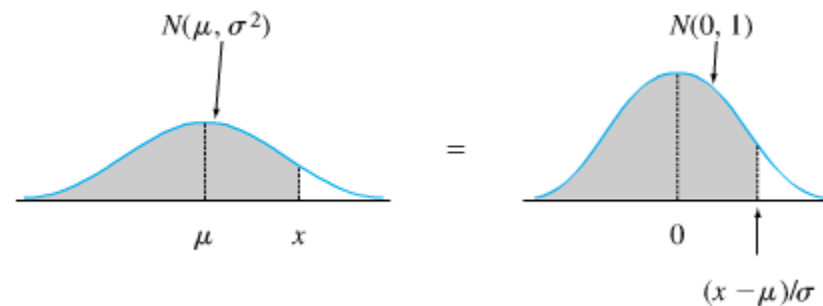
If  $X$  has a normal distribution with mean  $\mu$  and standard deviation  $\sigma$ , then

$$Z = \frac{X - \mu}{\sigma}$$

has a standard normal distribution. Thus

$$\begin{aligned} P(a \leq X \leq b) &= P\left(\frac{a - \mu}{\sigma} \leq Z \leq \frac{b - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{b - \mu}{\sigma}\right) - \Phi\left(\frac{a - \mu}{\sigma}\right) \end{aligned}$$

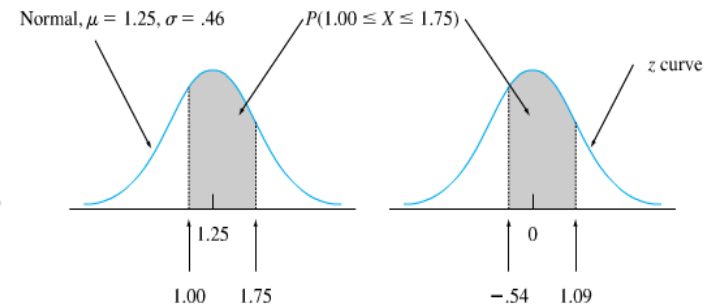
$$P(X \leq a) = \Phi\left(\frac{a - \mu}{\sigma}\right) \quad P(X \geq b) = 1 - \Phi\left(\frac{b - \mu}{\sigma}\right)$$



Example 4.16 Suppose that  $X$  is normally distributed with mean  $\mu = 1.25$  and standard deviation  $\sigma = 0.46$ , that is  $X$  is  $N(1.25, .2116)$ .

1. Find the probability  $P(1.00 \leq X \leq 1.75)$ .

$$\begin{aligned} P(1.00 \leq X \leq 1.75) &= P\left(\frac{1.00 - 1.25}{.46} \leq Z \leq \frac{1.75 - 1.25}{.46}\right) \\ &= P(-.54 \leq Z \leq 1.09) = \Phi(1.09) - \Phi(-.54) \\ &= .8621 - .2946 = .5675 \end{aligned}$$



2. Find the 90<sup>th</sup> percentile  $\eta(.90)$  of  $X$ , that is find a number  $c$  such that  $P(X \leq c) = 0.90$ .

The 90<sup>th</sup> percentile of  $Z$  is 1.28, which means that

$$P(Z \leq 1.28) = 0.90$$

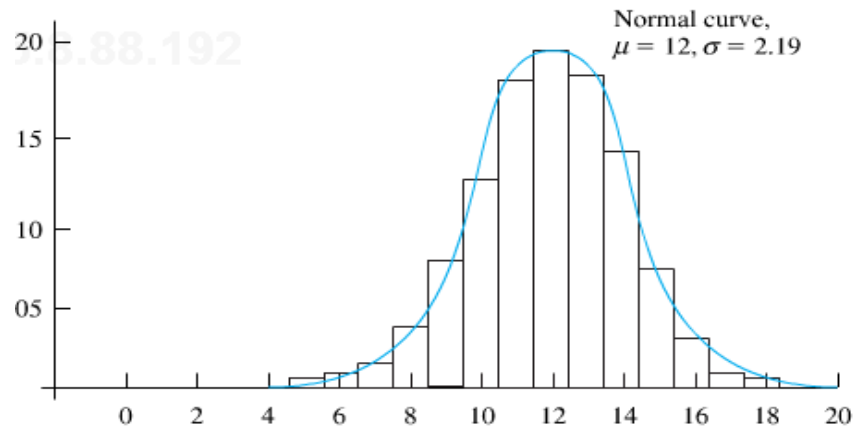
Since  $Z = \frac{X-1.25}{.46}$ , solving inequality  $\frac{X-1.25}{.46} \leq 1.28$  for  $X$  we obtain that  $P(X \leq 1.28 \times 0.46 + 1.25) = 0.90$ . Hence the 90<sup>th</sup> percentile  $c$  of  $X$  is  $c = 1.28 \times 0.46 + 1.25 = 1.84$

### TI-83:

1.  $P(1.00 \leq X \leq 1.75) = \text{normalcdf}(1,1.75,1.25,0.46) = 0.5681$
2. 90<sup>th</sup> percentile of  $X = \text{invNorm}(.9,1.25,0.46) = 1.84$

## Approximating Binomial Distribution:

For large  $n$  the cdf of binomial distribution with parameters  $n$  and  $p$  can be approximated by normal probabilities with  $\mu = np$  and  $\sigma^2 = np(1-p)$



Binomial probability histogram for  $n = 20, p = .6$  with normal approximation

### PROPOSITION

Let  $X$  be a binomial rv based on  $n$  trials with success probability  $p$ . Then if the binomial probability histogram is not too skewed,  $X$  has approximately a normal distribution with  $\mu = np$  and  $\sigma = \sqrt{npq}$ . In particular, for  $x =$  a possible value of  $X$ ,

$$\begin{aligned} P(X \leq x) = B(x, n, p) &\approx \left( \begin{array}{l} \text{area under the normal curve} \\ \text{to the left of } x + .5 \end{array} \right) \\ &= P(\tilde{X} \leq x + 0.5), \quad \text{where } \tilde{X} \text{ is } N(np, np(1-p)) \end{aligned}$$

In practice, the approximation is adequate provided that both  $np \geq 10$  and  $nq \geq 10$ , since there is then enough symmetry in the underlying binomial distribution.

### EXERCISES 4.3

Verify the Empirical Rule (below).

If the population distribution of a variable is (approximately) normal, then

1. Roughly 68% of the values are within 1 SD of the mean.
2. Roughly 95% of the values are within 2 SDs of the mean.
3. Roughly 99.7% of the values are within 3 SDs of the mean.

here SD = standard deviation =  $\sigma$

35. Suppose the diameter at breast height (in.) of trees of a certain type is normally distributed with  $\mu = 8.8$  and  $\sigma = 2.8$ , as suggested in the article “Simulating a Harvester-Forwarder Softwood Thinning” (*Forest Products J.*, May 1997: 36–41).
- a. What is the probability that the diameter of a randomly selected tree will be at least 10 in.? Will exceed 10 in.?
  - b. What is the probability that the diameter of a randomly selected tree will exceed 20 in.?
  - c. What is the probability that the diameter of a randomly selected tree will be between 5 and 10 in.?
  - d. What value  $c$  is such that the interval  $(8.8 - c, 8.8 + c)$  includes 98% of all diameter values?
  - e. If four trees are independently selected, what is the probability that at least one has a diameter exceeding 10 in.?

Answers: 0.3341, 0.0000316, 0.5785, 6.52, 0.8034



- 40.** The article “Monte Carlo Simulation—Tool for Better Understanding of LRFD” (*J. Structural Engr.*, 1993: 1586–1599) suggests that yield strength (ksi) for A36 grade steel is normally distributed with  $\mu = 43$  and  $\sigma = 4.5$ .
- What is the probability that yield strength is at most 40?  
Greater than 60?
  - What yield strength value separates the strongest 75% from the others?

Answers: a. 0.2525, 0.00008; b. 39.96

- 54.** Suppose that 10% of all steel shafts produced by a certain process are nonconforming but can be reworked (rather than having to be scrapped). Consider a random sample of 200 shafts, and let  $X$  denote the number among these that are nonconforming and can be reworked. What is the (approximate) probability that  $X$  is
- At most 30?
  - Less than 30?
  - Between 15 and 25 (inclusive)?

Answers: 0.993, 0.987, 0.792